

DO ASSET PRICES REFLECT FUNDAMENTALS? FRESHLY SQUEEZED EVIDENCE FROM THE OJ MARKET

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Abstract

The behavioral finance literature cites the frozen concentrated orange juice (FCOJ) futures market as a prominent example of the failure of prices to reflect fundamentals. This paper reexamines the relation between FCOJ futures returns and fundamentals, focusing primarily on temperature. We show that when theory clearly identifies the fundamental, i.e., at temperatures close to or below freezing, there is a close link between FCOJ prices and that fundamental. Using a simple, theoretically-motivated, nonlinear, state dependent model of the relation between FCOJ returns and temperature, we can explain approximately 50% of the return variation. This is important because while only 4.5% of the days in winter coincide with freezing temperatures, two-thirds of the entire winter return variability occurs on these days. Moreover, when theory suggests no such relation, i.e., at most temperature levels, we show empirically that none exists. The fact that there is no relation the majority of the time is good news for the theory and for market efficiency, not bad news. In terms of residual FCOJ return volatility, we also show that other fundamental information about supply, such as USDA production forecasts and news about Brazil production, generate significant return variation that is consistent with theoretical predictions. The fact that, even in the comparatively simple setting of the FCOJ market, it is easy to erroneously conclude that fundamentals have little explanatory power for returns serves as an important warning to researchers who attempt to interpret the evidence in markets where both fundamentals and their relation to prices are more complex.

1 Introduction

There has been considerable research that investigates whether asset prices reflect their fundamental values.¹ The general conclusion from this research is that the variability of returns is much greater than that implied by the assets' fundamentals, i.e., information about the asset's cash flows and expected returns. Specifically, regressions of asset returns on *ex post* information (relevant for pricing the assets) reveals little explanatory power.

Explanations for the empirically documented lack of a relation between asset returns and fundamental news can be found in the growing literature in behavioral finance. For example, DeLong, Shleifer, Summers, and Waldmann (1990, p.724-726) explicitly cite the above evidence as support for their noise trading model. It seems premature, however, to view this evidence as definitive. There is the alternative explanation that the relation between asset prices and fundamental information is incorrectly measured or not measured at all (e.g., omitted variables).² We provide supportive evidence favoring this latter view by analyzing a market where existing empirical results have been interpreted as being consistent with market irrationality, and to illustrate how looking at the data in a new way can potentially overturn this interpretation. By necessity, we choose an example where at least some of the relevant fundamental information is easy to identify, and where theory provides strong intuition about the functional form of the relation between fundamentals and returns. The fact that, even in this simple setting, it is easy to erroneously conclude that fundamentals have little explanatory power for returns is an important warning to researchers who attempt to interpret the evidence in markets where both fundamentals and their relation to prices are more complex.

Our example comes from one of the forerunners of the excess volatility literature—Roll's (1984) seminal paper on frozen concentrate orange juice (FCOJ) futures prices and weather. His paper is unique and clever in that the fundamental information appears to be identifiable and exogenous. A critical aspect of this market is that the relevant oranges for FCOJ are produced in a relatively

¹For example, see Shiller (1981), Meese and Rogoff (1983), Black (1986), Campbell and Shiller (1987), Frankel and Meese (1987), Roll (1988), Cutler, Poterba and Summers (1989), Lev (1989), and Mitchell and Mulherin (1994) to name just a few.

²Moreover, the research design of some of these previous studies have been put into question, examples of which include Kleidon's (1986) analysis of Shiller's (1981) excess volatility paper, or, more recently, Fama's (1998) criticism of so-called anomalies in the corporate event study literature.

small region of Florida. Thus, as Roll (1984) reasonably argues, the weather in this region should be a primary determinant of supply and thus of futures prices. Roll (1984, p.876) states that “weather is the most identifiable factor influencing FCOJ returns”; however, overall, he finds little explanatory power in the relation between FCOJ futures prices and weather. In his conclusion, Roll (1984, p.879) states that

...weather surprises explain only a small fraction of the observed variability in futures prices. The importance of weather is confirmed by the fact that it is the most frequent topic of stories concerning oranges in the financial press and by the ancillary fact that other topics are associated with even less price variability than is weather... There is a large amount of inexplicable price volatility.

Weather’s apparent lack of explanatory power for FCOJ returns has become a popular citation in the behavioral literature precisely because the fundamental information seems to be so clear. However, it is important to note that while Roll’s findings have subsequently been adopted by the excess volatility literature, at the time, the paper was seen in many ways as being supportive of market efficiency in that it also focused on how prices rationally incorporate all the information in weather forecasts. Moreover, while Roll identifies the puzzling lack of explanatory power, he does not argue that this evidence implies market irrationality. Nevertheless, in recent surveys of behavioral finance

- Shleifer (2000, p.20), in a discussion about empirical challenges to efficient markets, asks “what about the basic proposition that stock prices do not react to non-information? Here again there has been much work, but three types of findings stand out... many sharp moves in stock prices do not appear to accompany significant news... A similar conclusion has been reached in two striking studies by Richard Roll (1984, 1988)...He finds that, although news about weather helps determine future price movements, they account for a relatively small share of these movements”;
- Hirshleifer (2001, p.1560) writes in a section covering mispricing effects that “little of stock price or orange futures price variability has been explained empirically by relevant public news”;

- Ross (2001, p.11) argues that “the regressions themselves have low power even when run with the return as the dependent variable and past returns and weather forecasts as well as contemporaneous weather changes. The R^2 of these regressions is on the order of 1 or 2%. In other words, while there is no evidence that prices do not fully incorporate past information, we are pretty much at a loss to say why they move at all! If it’s not weather - which is obviously the biggest determinant of supply changes - then what could it be?”;
- Daniel, Hirshleifer and Teoh (2002, p.172) state “only a small fraction of stock prices or orange juice futures prices has been explained by the arrival of relevant public news. Roll (1984) finds that the volatility of OJ futures prices was hard to explain by news about the weather”.

In this paper, we show that this literature has misinterpreted the data by ignoring the state dependence inherent in the structural relation between FCOJ returns and their underlying fundamentals. In particular, the identified fundamental, i.e., weather surprises, should theoretically impact FCOJ futures returns only in one state of nature, namely to the extent this surprise provides the market with some new information about the probability and severity of a freeze. The implication of this is that most weather surprises should have no impact at all. For example, an unexpected 30 degree drop in temperature from 80 to 50 degrees in late December might affect the tourism industry in Orlando, but it will have no impact on the FCOJ market. In contrast, a much smaller unexpected 5 degree drop from 30 to 25 degrees will have devastating effects on the Florida orange crop.

We document that in most circumstances, as theory would predict, the R^2 is zero between FCOJ futures returns and weather changes. However, around freezing temperatures, the only state of nature that theoretically matters, we estimate R^2 s on the order of 50% using a variety of (albeit rudimentary) models and information. This difference in R^2 is nontrivial. In fact, over one-half of all variation of FCOJ returns occurs in the winter season. Of this variation, two-thirds occurs on just 4% of the days. Not surprisingly, these days are not random and are instead related to freezing temperatures. There is only conclusion that can be drawn from these results - there is absolutely no puzzle concerning the relation between FCOJ prices and the weather.

There is, of course, variation of FCOJ returns in other seasons when there is little or no information about the possibility of freezing temperatures. However, the fact that FCOJ prices

vary throughout the year should come as no surprise to agricultural economists; this is true of all commodity prices (see, for example, Fleming, Kirby and Ostdiek (2002)). Whether the remaining volatility can be explained by information about demand for OJ, or futures-specific issues such as the microstructure of the market or the commodity's convenience yield, is difficult to determine. Nevertheless, we focus on other non-weather related news about supply-side shocks to the FCOJ market and identify three important sources: (i) USDA production forecasts, (ii) resolution of production uncertainty, and (iii) production from areas outside of Orlando, the majority of which comes from Brazil. Theory suggests that all of these sources should generate news that is important for FCOJ prices, and the evidence confirms this. For example, each year in mid-October the United States Department of Agriculture (USDA) provides its initial forecast of orange production for the season. Over 40% of total fall FCOJ futures return variance occurs on that day. This fact alone suggests that the market does not ignore fundamental information about production independent of the weather effect.³ Together, the October production forecast days and freezing temperature days account for 1.5% of the days in our sample. On these days 41% of total FCOJ return variance is observed. This is apparently not a market in which investors are ignoring public fundamental information.

The paper is organized as follows. In Section 2, we provide the theory for how FCOJ prices should be related to supply shocks, and, in particular, weather. Of particular interest, we shed light on why the literature has misinterpreted the existing findings. Section 3 describes the data and the basic stylized facts about weather, production and FCOJ futures prices. This section provides strong circumstantial evidence that these three variables are linked. In Section 4, we then test these links more formally, and, in general, confirm the theoretical relation empirically. In particular, there is close relation between FCOJ prices and temperature when the theory suggests there should be one. Section 5 explores the fact that there is remaining volatility in FCOJ prices not explained by the weather. In particular, we explore the effect of other factors on FCOJ prices. Section 6 concludes.

³Bauer and Orazem (1994) examine USDA forecasts, showing that forecast surprises provide information that is relevant for FCOJ pricing. They report R^2 s for the initial October forecast day that are similar in magnitude to those we report for freeze days.

2 Theory of FCOJ and Weather

There is a growing literature in finance arguing that asset prices drift away from their underlying fundamentals due to irrational agents who ignore these fundamentals. This literature is motivated, at least in part, by empirical evidence documenting a disconnect between fundamentals and asset prices. For example, the empirical literature reports low R^2 s from linear regressions of returns on unexpected changes in the fundamentals. However, these regressions measure the linear, state independent, relation between innovations to the fundamentals and asset returns, and attempting to link asset returns to changing fundamentals in such an atheoretical framework can be problematic. Asset pricing models often result in a complex structural relation between fundamentals and prices; therefore, linear regressions ignore state dependence and nonlinearity that occur quite naturally.

This paper addresses the problematic nature of drawing economically meaningful conclusions from these types of empirical models in the context of the FCOJ market. The FCOJ market is a natural choice because it has generated considerable attention, not least because it seems to provide such a strong case for irrationality. What are the fundamental determinants of prices in this market? First, and foremost, there are short-run shocks to the Florida-based production of oranges,⁴ which are primarily attributable to the weather. Second, due to worldwide production of oranges suitable for FCOJ, unexpected changes in the ability of distributors to shift away from Florida oranges to other sources also affects returns. In particular, Brazil is the largest exporter of oranges worldwide and thus news about production in Brazil is relevant for the price of FCOJ. Finally, there could be short-run demand shifts either to or away from FCOJ. Competing products include other citrus juices, such as grapefruit juice, lemonade or freshly squeezed orange juice, and other fruit juices such as apple juice. The fact that companies spend considerable amounts marketing FCOJ products provides anecdotal evidence that demand matters.

Thus, relative to other markets, the FCOJ market appears to be less complicated. Of equal importance, the primary fundamental factor, namely weather, is clearly exogenous. No one believes FCOJ prices have any affect on weather patterns in Orlando, Florida. It is this combination of features that should provide a clean environment for examining the relation between asset prices

⁴In fact, because storage costs of oranges and, in particular, FCOJ are high, only a fraction of the juice gets carried forward from year-to-year. Therefore, long-run shifts in either demand or supply, are expected to have little impact on FCOJ prices.

and fundamentals.

Outside of the fundamental determinants of FCOJ prices, there may be additional factors that affect FCOJ returns. However, since it is not our goal in this paper to explain *all* the movement in prices, but rather to demonstrate that relevant information about fundamentals gets incorporated into prices, we ignore a lot of return-relevant information. In particular, we ignore well known factors that affect returns including microstructure effects (e.g., liquidity and the bid-ask spread), the role of stochastic convenience yields, variations in storage costs, stochastic interest rates, delivery costs and so forth, which are covered elsewhere in the literature.

There are several reasons why a simple linear model cannot provide a sufficiently rich framework to address the question of whether asset returns incorporate information on fundamentals efficiently in the context of the FCOJ market. Specifically,

- The relation between FCOJ returns and weather surprises depends on the existing state.
- There is a nonlinear relation between FCOJ returns and weather.
- The realization of the expected temperature (i.e., a zero forecast error) is itself news.

The following three subsections elaborate on these issues.

2.1 State Dependence

Roll (1984) performs three analyses by looking at the relation between FCOJ futures returns and (i) temperature surprises, (ii) rainfall forecast errors, and (iii) a variety of explanatory variables including temperatures, exchange rate changes, and stock returns. Proponents of market irrationality focus their attention on some of the results related to (i). With respect to temperature surprises (the difference between realized temperatures and forecasts), Roll documents a statistically significant relation, but finds R^2 s between 1% and 4%. Returns are informationally efficient in that forecast errors are not related to future returns, which is arguably the focus of his analysis, but the explanatory power is minimal. Based on these results, Roll concludes that “the small predictive power for temperature and rainfall seems to imply that influences other than weather are affecting OJ returns” (p.875). It is empirical evidence and statements such as these that have been interpreted by other researchers as indicating a lack of explanatory power for fundamentals in the FCOJ futures market.

However, the relation between returns and weather surprises depends critically on the existing state. That is, the conventional view in the FCOJ market is that it is only freezing temperatures that produce supply shocks. Thus, theory implies that the weather surprise is only relevant to the extent it tells us something about the change in likelihood and magnitude of a freeze.

There are several obvious ways to incorporate state dependence. First, freezes are seasonal – fall, spring and summer months are unlikely to produce freezing weather in Florida.⁵ Second, a temperature surprise, or change in temperature, will only matter around freezing temperatures. That is, a 10-degree surprise at 70 degrees has very different implications for the likelihood and severity of a freeze than it does at a temperature of 35 degrees. Even during winter, at most temperature levels, theory tells us that there is no relation between FCOJ prices and temperature. Third, the distribution of the freeze-related supply shock will depend on whether there has already been a freeze during the same growing season. For example, if a freeze has a substantial impact on orange production, then a second freeze may have less impact to the extent that the damage is already done. Of course, this depends on the relative severity of the freezes, the timing of the harvests, etc. In addition, conditional on a freeze, the uncertainty regarding its effects on production can be quite high until the damage is surveyed. Any analysis of R^2 s that does not take these state dependencies into account is flawed.⁶

2.2 Nonlinearity

A second obvious improvement to the linear specification is to allow the unexpected change in the fundamental to be nonlinearly related to the asset return (controlling for the state dependence

⁵Roll attempts to adjust for the importance of freezes by giving greater weight to observations in winter, but this correction is inadequate to capture the degree of state dependence in the data. In particular, we will argue that the weight on non-winter data must be zero by theory.

⁶In this paper, we do not look at the relation between the FCOJ market and rainfall. Similar to temperature surprises, however, Roll runs a regression on rainfall forecast errors and finds no discernible relation to returns. With respect to rainfall, Roll acknowledges that its effect “on the crop is much less obvious than the effect of temperature” (p.873). For example, either too little or too much rainfall can have a negative impact on OJ production and quality. Moreover, only extended periods of too little or too much rain are likely to have a significant effect, and the magnitude of the effect will depend on the timing relative to the growing season. Thus, the analysis with rainfall is also suspect in terms of interpreting R^2 s due to the presence of similar state dependence and nonlinearity as observed with temperature.

as discussed above). In practice, the severity of a freeze depends on two factors: (i) the level of the temperature, and (ii) the duration of a freeze. Table 1 provides a brief description of these relevant factors. In particular, the severity of a freeze is not linear in temperature, as 30-31 degree temperatures have relatively mild effects compared to temperatures in the 27-29 degree range. In the latter instance, there is severe damage done to both the leaves and fruit, which dramatically affects production. Thus, at freezing levels, as temperatures drop just a degree or two, the supply shock gets worse at an accelerating rate. On the other hand, at levels much below freezing, another degree or two will not make a difference. The effect of ignoring the nonlinearity will again be to lower the R^2 s in the regression analysis.

2.3 News When Forecasts are Correct

Implicit in a regression of asset returns on forecast errors is the assumption that a zero forecast error (i.e., the realization of the expectation) should have no effect. While this holds for a linear relation between prices and fundamentals, it does not hold generally. Consider a scenario under which the market forecasts a freezing temperature but with some uncertainty. In other words, the expected minimum temperature tomorrow is below freezing, but, given the difficulties inherent in weather forecasting, there is a significant probability the temperature will be higher or lower. The full distribution of possible future temperatures is incorporated in the price of FCOJ futures today. In particular, today's price reflects the probability that there will be no freeze tomorrow, no supply shock and therefore no impact on prices. When the freeze is realized (and let us suppose at the forecasted minimum temperature so that the "weather surprise" is zero), there is an obvious shift in the likelihood of a freeze. Invariably, returns will increase in this instance even though the forecast error is zero. Of course, ignoring this fact leads to lower R^2 s for FCOJ returns and in other cases in which asset returns are linked to "surprises" in the fundamentals, as long as there is state dependence or nonlinearity in the relation.

While Roll (1984) does not explicitly address the importance of this possibility in his analysis, he does report results from a regression of returns on a temperature freeze variable and seven other variables including exchange rates, oil prices and stock returns (see Roll (1984), Table 10 in a section entitled "Nonweather Influences on OJ Prices"). Though we take issue with the exact form of the specification, the identification of the freeze variable is very close to identification of

the relevant fundamental. This is because the freeze variable captures both the state dependence described in Section 2.1 above and the realization versus expectation issue described here. Not surprisingly, the explanatory power of this regression is greater, although the R^2 is still only 6.7%. This low R^2 derives from (i) treating all observations the same (e.g., running the regression across all seasons), (ii) not incorporating temperature forecasts, (iii) ignoring nonlinearities between weather and returns, and (iv) ignoring limit price moves (in contrast to Roll's earlier tables). Thus, it is perhaps not surprising that the existing behavioral literature has drawn the wrong inference from these results.

3 Data

Futures contracts in frozen concentrated orange juice (FCOJ) have traded on the New York Cotton Exchange since September of 1967. At any given time, there are usually nine to eleven contracts outstanding with expiration schedules every second month (January, March, May, etc.) and with at least two January months listed at all times. The contract is for 15,000 pounds of frozen concentrated orange juice, which represents about 2,400 ninety-pound field boxes of oranges, with specific requirements for color, favor, and defects. Due to these requirements, 95% of the total U.S. processed orange production takes place in central Florida in and around the Orlando area.⁷ There are two types of oranges produced in the Orlando area with the main distinction being the harvest period, namely early and mid-season (EM) varieties and Valencia oranges, which get harvested from November through March (though primarily in January) and April through June, respectively.

The geographic concentration in orange juice production is somewhat unique for agricultural commodities, and allows us to study the interaction between asset returns (i.e., FCOJ futures prices) and an important exogenous variable (i.e., Orlando weather). *Ex ante* there are strong reasons why weather should be an important variable for FCOJ prices and why Roll's (1984) result is such a puzzle. In particular, freezes can have devastating effects on orange production. Attaway (1997)

⁷California is also a large producer of oranges; however, oranges produced in that region are not suitable for FCOJ. Internationally, Brazil is, along with the U.S., the largest producer of oranges used for FCOJ. While Brazil primarily exports these oranges to countries other than the U.S., the majority of Florida production is consumed domestically. Though Brazilian oranges are subject to significant tariffs, they can provide a close substitute to Florida oranges. The issue of Brazil's production and how it affects the FCOJ market is discussed in Section 5.2.

cites several freezes over the past 170 years that have had a major impact on fruit production. A full list of the impact freezes is provided in Table 2. For example, during the freeze of February 7-9, 1835, temperatures dropped to as low as 11 degrees in northern Florida, destroying almost all existing orange trees. In fact, prior to this date, a number of more northern states (such as South Carolina and Georgia) produced oranges. The freeze of 1835 essentially forever ended the desire to produce oranges in those states. In the decade from 1894 to 1905, Florida was hit by ten freezes, one of which (February 8, 1895 when the temperature reached 17 degrees) led to almost all the orange crop being destroyed. In 1893-94, 5.06 million boxes of oranges were produced; by 1895-96, production had fallen to 0.147 million boxes (Attaway (1997)). The industry did not fully recover until 1909-10. Of interest to this paper, Table 2 shows that a number of freezes occurred during our sample period, that is, post 1967.

We collected data from September 1967 to December 1998 on several series: (i) the minimum temperature in Orlando, (ii) a limited amount of minimum temperature forecasts from the National Weather Service (primarily for winter months and for only some of the years within our sample), (iii) USDA production forecasts, and (iv) FCOJ futures prices, volume and open interest on every contract. We also collected other relevant data series such as Brazilian orange production and FCOJ exports to the U.S., and news events associated with FCOJ. We describe the data in detail as it becomes relevant throughout the paper.

3.1 Orlando Weather

The most important fundamental for FCOJ prices is the presence of a production shock to a given year's orange crop. The largest factor in this regard is weather, and, in particular, winter freezes. As mentioned above, winter freezes can severely impact orange production, thus raising the prices of oranges in the commodity market.

Figure 1 graphs the minimum temperature in Orlando, Florida from 1967 through the end of 1998. Over 80% of the winter years include one or more days with temperatures close to or below freezing during periods relevant for orange juice production. Specifically, within our sample period of thirty-one winters, only six winters can be strictly described as completely freeze-free winters. Twenty-five winters experienced one or more nights when the minimum temperature dipped to 32 degrees or below. Approximately 50% of the winters have two to three nights of freeze, and the

probability of four or more freeze nights is approximately 25%. Thus, there is clearly a considerable amount of price-relevant information during the winter season.

However, it is important to point out that not every freeze has material consequences for the orange crop. As mentioned above, Attaway (1997) provides a detailed description of the history of impact freezes that happened in Florida. While Table 2 documents the actual dates of these freezes, recall that Table 1 describes the critical conditions for freezes to have a substantial effect on production. These conditions are essentially the temperature itself and the duration of this temperature level.⁸ For example, if the temperature falls below 28 degrees for six hours or more, then, without intervention, oranges will suffer extensive damage, hence affecting production of FCOJ. With respect to our particular sample period, according to Attaway (1997), there are eleven relevant freeze seasons, some of which have multiple freezes, from 1967 to 1998.⁹ The identification of these freezes will form the basis for our comparison between freeze and non-freeze years.

3.2 Production

Every October, the United States Department of Agriculture (USDA) provides forecasts of the upcoming season's orange production. This forecast is conditioned on a non-freeze season, that is, the USDA does not take into account the possibility of a freeze when forming its forecast.

Table 3 documents the percentage difference between the October USDA forecast and final production, which is generally reported in August of the following year, for all the years in our 30-year sample. Freeze years, as classified by Attaway (1997), are noted, and we also report the average across the 11 freeze years and 19 non-freeze years. Several observations are of interest. First, the USDA's forecasts in non-freeze years are nearly unbiased, with only a 1.3% difference on average between the forecast and the realization. However, there is substantial variation in forecast accuracy, with underestimates and overestimates as large as 6.2% and 8.1%, respectively. Second, and most important, there is a considerable reduction in orange production in freeze years, with

⁸Of course, the magnitude of a freeze alone is not sufficient to explain the impact on prices. Also important are coincidental effects on the determinants of the supply of Florida FCOJ substitutes such as imported FCOJ from Brazil. Examples of these effects are whether there was a drought in the Sao Paulo area of Brazil, tariff rates, and the costs of transportation.

⁹In a separate study by Hebert (1993), similar freezes are also identified using a different freeze identification methodology.

12.7% less production than forecast. Again, there is substantial variation, with a maximum decline of 30.5%. If the demand for oranges is downward sloping and there is no perfect substitute for Florida oranges (at the prevailing price), then the spot price of oranges should rise in response to a decline in supply. Given the economic laws of supply and demand, therefore, futures prices on FCOJ should move in the same direction.

3.3 Futures Prices

In order to investigate the effect of freezes on prices, we obtained daily closing prices of FCOJ futures from September 1967 to August 1998. On each trading day prices of the three near maturity contracts were collected from Bridge/CRB and Datastream. We also collected volume and open interest data. There are a number of ways to generate a continuous return series from the three price series. In general, we attempt to get the most accurate proxy for price changes in spot FCOJ prices. We need to take into consideration two important characteristics of futures' prices that may affect the accuracy of calculated returns – liquidity and limit days. First, it is well known that near maturity contracts are the most liquid, a fact that we verified with the volume and open interest data. However, liquidity diminishes rapidly as a contract comes close to expiration, particularly in the expiration month. These factors were considered in constructing the continuous return series. A second issue is that price moves on futures contracts are limited by the exchange. The magnitude of these price limits and the contracts to which they applied change over the course of our sample, but the general impact of price limits is to prevent prices from fully incorporating information on days with important information releases. We adjust for this phenomenon in the standard way by aggregating returns over consecutive days with price moves that hit either the lower or upper limit. For example, when a severe freeze occurs, the price may hit the upper limit for 4 consecutive days. In this case, we aggregate returns over these 4 days plus the next day, which is a no-limit day, and use this as the 1- day return. The precise algorithm we use to construct the spliced data series is described in the appendix.

Given this continuous daily return series, Table 3 provides the cumulative return over the winter season (December through February) for each year and the average of these returns for freeze and non-freeze years. The average return is almost 19% higher in freeze years than in non-freeze years, namely 12.7% versus - 6.1%, although there is substantial variation in returns on an annual basis.

This suggests that, on average, futures prices are indeed higher in years in which production has been impacted negatively.

It is also worthwhile documenting some additional stylized facts that provide substantial circumstantial evidence of an important relation between FCOJ futures prices and temperature. Table 4, Panel A reports the volatility of daily FCOJ returns in different seasons. The variance of FCOJ futures returns in winter months is four times that of returns in spring and summer months and three times higher than in the fall season. This is not surprising as freezes only occur in winter months. A slightly different way of viewing the data is to note that more than 50% of the total variance of returns is accounted for by the winter season. If the mean returns are the same across different categories, as they are approximately across seasons, this variance decomposition can be calculated simply by multiplying the daily variance by the number of observations in any season.

We also divided the winter season into two separate periods: pre-freeze and post-freeze. The pre-freeze period includes days up to and including the first freeze of the season in a freeze season and all the days in a non-freeze season; the post-freeze period includes the days subsequent to a freeze in a freeze season. Note that the post-freeze period can include a second freeze, and, in fact, two seasons had multiple freezes (see Table 2).

Interestingly, the variability in winter months is greater pre-freeze than post-freeze in spite of the fact that freeze frequencies are similar in the two periods. Of course, this phenomenon may be due to the fact that the second freeze in a season happened to be less severe in our sample. Alternatively, the second freeze in a season may be less important because damage has already been done to the orange crop. Due to uncertainty surrounding the impact of multiple freezes, the remainder of the paper focuses on the period up to and including the first freeze of the season, i.e., the pre-freeze period.

The fact that winter months produce greater variation in futures prices does not in itself suggest a relation between temperature and FCOJ prices. As shown in Tables 1 and 2, the main impact on production comes from a particular event, namely a freeze. Thus, it is the change in the likelihood of a freeze that should move FCOJ futures prices. Table 4, Panel B presents the volatility of FCOJ futures prices in winter months conditional on various contemporaneous minimum temperature realizations, namely 35 degrees and below, 36-40 degrees, 41-45 degrees and 46 degrees and above.

As can be seen from the table, at or around freezing temperatures (i.e., 35 and below), the

daily standard deviation of FCOJ returns is 11.81% compared to 1.79%, 2.08% and 2.16% for the other temperature buckets. That is, there is over 50 times greater variance in FCOJ futures returns near freezing temperatures than when temperatures are warmer. As a result, in spite of the fact that there are only 77 observations at these low temperatures (5.2% of the winter pre-freeze sample and 1.0% of the total sample), these observations account for almost 70% of the variance in the winter pre-freeze period and over one third of the total variance.¹⁰ This result strongly suggests that freezes have a substantial impact of FCOJ futures pricing.

While there are many factors that can affect the FCOJ futures price, we have documented several important stylized facts with respect to FCOJ futures price changes and supply shocks due to temperature (freezes). First, over 80% of the years have freezing weather in Orlando, Florida (Figure 1). Moreover, 40% of these years are considered freeze years (Table 2); that is, years in which the temperature was low enough for a long enough time to be relevant for orange production. Second, adjusting for expectations about production, there is a 14% difference in production levels for freeze versus non-freeze years (Table 3). Consistent with agricultural theory, this result confirms that freezes have substantial impacts on U.S.-based production. Third, with respect to financial markets, the distribution of futures returns differs in freeze versus non-freeze years. Specifically, returns are on average 19% higher in freeze years (Table 3). This evidence is consistent with financial markets relating asset prices to the fundamentals, i.e., futures prices are much higher in years with negative shocks to production. Fourth, the majority of the well-documented variation in futures returns occurs in the temperature region one would expect, namely around freezing temperatures (Table 4B). For example, while only 5.2% of the winter pre-freeze days are in this region, 69.4% of the variance in futures returns occurs on these days.

4 FCOJ Prices and Temperature: Empirical Results

Given these four facts, it seems surprising that one of the major results in the excess volatility literature is the lack of explanatory power in the relation between FCOJ futures and temperature. In a rational market the FCOJ futures return over a given period should reflect changes in expectations

¹⁰These variance decomposition calculations are slightly more complicated than those across seasons because the mean returns differ across temperature levels. Consequently, sums of squared deviations are calculated using the overall mean.

about the level of short-run production, i.e., the supply shocks over that period. With respect to temperature, these shocks correspond to changes in the market’s perceived likelihood and severity of a freeze.

Perhaps the literature has found little or no relation because the change in likelihood of a freeze is a difficult variable for the econometrician to measure. Clearly, there is no theoretical reason to believe that the relation between FCOJ futures returns and temperature surprises fits into a linear framework. Section 2 above argues that there is considerable reason to believe this is not the case. We identify several important factors such as (i) the existing state of nature (e.g., season, pre-freeze or post-freeze, current freeze probability), (ii) the realization versus expectation in terms of measuring the change in the freeze’s probability distribution, and (iii) the nonlinear nature of freeze severity based on both the temperature level and the freeze duration. In the next subsection, we try and relate these factors more closely to the data by explicitly incorporating some of these ideas into an empirical analysis of futures prices and temperature.

4.1 A First Look

Consider the linear regression of FCOJ returns on temperature surprises:¹¹

$$R_t = \alpha + \beta(W_t - E_{t-1}[W_t]) + \epsilon_t,$$

where R_t is the close-to-close return, W_t is the contemporaneous realized minimum temperature, and $E_{t-1}[\cdot]$ is its forecast the previous period. As discussed in Section 2, this specification of FCOJ returns is problematic.

A better alternative would be to project FCOJ returns on the changes in the market’s perception of the likelihood and severity of a freeze. While this variable is unobservable, there is one instance in which this is straightforward to do – when a freeze actually occurs. In this case, the likelihood is probability one and the severity of the freeze maps to the temperature level. Even in this simplified case, the temperature level is just one of two pieces of information. The other piece is the length of time at each temperature level, which we discuss in the next subsection. Of course, market participants have access to greater information such as verbal assessments of the probable freeze damage, the particular timing of the freeze as it relates to the harvesting period, etc.

¹¹Roll (1984) runs this regression, although the dependent variable and independent variables are switched. The estimation results form the basis for the conclusions in the behavioral literature described in the introduction.

If one assumes that the conditional probability of a freeze is constant, then the realization of a freeze and its corresponding temperature level will measure the “freeze surprise” and therefore the supply shock (albeit functionally). There are strong reasons to question the assumption that the probability of a freeze is constant. At the very least, the probability will vary with the time of the year. More important, there is information provided to the marketplace about the likelihood and severity of freezes throughout the year. For example, the U.S. National Weather Service provides 12, 24 and 72 hour forecasts that the market should incorporate in its assessment. These issues are addressed in the next subsection in some detail.

The functional form of the return-temperature relation is unknown, but is certainly nonlinear. We assume that such a relation exists, and attempt to characterize the function empirically. We investigate several different specifications for this relation. The relation between FCOJ returns and weather is theoretically only relevant around freezing temperatures; thus, we can ignore the seasons outside of winter. Further, the functional relation between these variables may depend on whether a freeze has occurred, so we condition on pre-freeze winters.

Consider the following linear model:

$$R_t = \alpha + \beta W_t + \epsilon_t,$$

where the return and the minimum temperature are aligned in time. We consider this the benchmark model, and it is equivalent to a regression on temperature surprises if there is no forecasting power. The clear problem with the linear model is that it imposes the same relation between returns and temperature at temperatures both above and below freezing.

In contrast, an alternative specification, the piecewise linear model, accounts for the state dependence by estimating a number of linear relations, each for a different temperature bucket,

$$R_t = \sum_{n=1}^N D_{nt}(\alpha_n + \beta_n W_t) + \epsilon_t,$$

where the dummy variable, D_{nt} , equals one for $W_t \in (W_{min}^n, W_{max}^n]$ and zero otherwise. Each return-temperature pair is placed in a temperature bucket and is used in estimation of the relevant coefficients α_n and β_n . In theory, this model provides complete flexibility. Given enough data the econometrician can specify a large number of small buckets, effectively estimating the functional form of the relationship piece-by-piece. Note that we do not impose continuity of the return-weather function, that is, there is no the restriction that the lines connect at the bucket borders.

Under these assumptions, the most general model estimates an unspecified functional form linking returns to temperature:

$$R_t = f(W_t) + \epsilon_t.$$

In this paper, we estimate this functional form using the kernel regression methodology described in Ullah (1988) and others. The kernel estimation provides a standard nonparametric regression of returns on realized temperatures based on nonparametric density estimation. The econometrician chooses the kernel function and the bandwidth or degree of smoothing. The results are insensitive to the choice of kernel function and we use a normal density. The bandwidth is chosen via cross-validation. Specifically, we estimate the fitted returns at each temperature observation in the data using all the observations except that specific point and choose the bandwidth that minimizes the resulting mean squared error.

One could also impose a functional form. To the extent that the theory implies only news about freezing weather matters, a natural specification given our constant freeze probability assumption is

$$R_t = \alpha + \beta_1 \text{Max}(0, W^* - W_t) + \beta_2 [(\text{Max}(0, W^* - W_t))]^2 + \epsilon_t,$$

FCOJ returns are a second order polynomial in the regressor $\text{Max}(0, W^* - W_t)$, where W^* represents the critical temperature threshold. In our analysis, we use 32 degrees as this threshold, although the results are insensitive to any reasonable choice. To the extent FCOJ prices are not affected by non-freezing temperatures, the regressor is zero. Then, as the temperature gets below a certain level, the regressor rises quickly as the temperature falls. Thus, this model incorporates the nonlinearity implied by the theory of freeze impacts described in Table 1. As pointed out in Section 2.3, Roll (1984) also uses the variable $\text{Max}(0, 32 - W_t)$, although he does not include the squared term and he includes a variety of other variables including exchange rates, oil prices and stock returns.

Table 5 presents the results for these four different specifications estimated using the winter pre-freeze sample. The results are in stark contrast to the existing literature's view of the relation between FCOJ futures prices and temperature (e.g., Shleifer (2000) and Hirshleifer (2001)). While temperature is statistically significant in the linear model, it is clear that the evidence derives almost entirely from low temperatures. The piecewise linear model shows that weather only matters at low temperatures. Of course, this is precisely what the theory would predict. That is, in terms of

the weather, it is only information about a freeze, in this case, the realization of one, that matters.

The R^2 s jump from 1.5% for the linear model to as high as 33.3% and 33.6% for the nonlinear model and kernel nonparametric models, respectively. Thus, across the entire winter pre-freeze period, over one-third of the variation can now be explained. To get a finer partition of the explanatory power of weather, Figure 2 breaks up the R^2 s into different temperature buckets. This partitioning is important because the theory suggests that weather should only matter at low temperatures. The results confirm this theoretical prediction, with the more general models capturing close to 50% of the return variation for temperatures of 35 degrees and below. In fact, outside of the low temperature ranges, there is almost no relation between weather and FCOJ futures returns.

This breakdown of R^2 s for different temperature ranges requires some discussion because it is not a standard technique. First, note that we are not snooping the data by conditioning on large returns (the dependent variable). Instead, we are conditioning on the temperature (the independent/exogenous variable) that should, in theory, lead to large returns. Thus the R^2 s we report are neither spurious nor overstated; they accurately reflect the true explanatory power (asymptotically) in this region of the data. Second, this kind of analysis is only interesting in a nonlinear/state dependent context. In the case of a standard univariate linear regression model with homoscedasticity, the R^2 s increase as the deviation of the independent variable from its mean increases. In the general nonlinear case this is no longer true, and the R^2 s can take on any pattern conditional on the level of the independent variable. In fact, in our sample, conditioning on large deviations above the mean (i.e., the high temperature region) produces R^2 s of zero. The point of the analysis is to see if the pattern in the data is consistent with the theoretical predictions and to examine the magnitudes in the regions when the fundamental should matter.

Note that the linear model produces negative R^2 s for the three higher temperature ranges. A negative R^2 is possible because we estimate the model on all the data and then calculate the R^2 for specific groups of observations sorted by temperature. In theory, the other globally estimated models (the nonlinear and nonparametric models) could also produce negative R^2 s in some regions, but the actual R^2 s are positive but close to zero.

Given an R^2 of close to 50% in the temperature range that accounts for about one third of the total variance of returns (see Table 4), we would expect the regression model to have an R^2 of 16%

when estimated over the full sample. This prediction is confirmed by the results reported at the bottom of Table 5. Why is this number substantially higher than the 6.7% R^2 reported by Roll (Table 10, p.877)? There are two contributing factors—the absence of the squared term and the fact that Roll does not aggregate limit moves for this regression (although he does so elsewhere in his analysis). The former is responsible for about 2% of the difference, as shown in the final regression in Table 5, while the latter presumably accounts for the remainder.

The high R^2 s documented above demonstrate that FCOJ futures prices do react significantly, and in the correct direction, to fundamental information about freezes. However, it is still possible that prices are either under- or over-reacting to this news. Without a formal structural model relating freezes to supply shocks and thus price movements, it is hard to know exactly how large returns should be when freezes occur. Nevertheless, the time series properties of returns do provide important information about the potential existence of under- or over-reaction. For example, under-reaction to a news event implies a positive serial correlation in returns (at some horizon) as prices eventually reflect the fundamental information, and thus future price movements must, on average, be in the same direction as the price movement at the time of the event. Similarly, over-reaction implies negative serial correlation in returns. It is difficult to know at what horizon these effects might exist, but since substantial information about the effect of a freeze on production comes out within a relatively short period of time, it seems logical to examine returns over the month subsequent to a news event. There are also a number of possible ways to look for under- or over-reaction, e.g., examining autocorrelations of returns directly or looking at variance ratios. For simplicity we examine both autocorrelations and mean returns conditional on significant events.

Table 4, Panels A reports the first order autocorrelation and the average of the autocorrelations at lags 2 through 20 for daily returns over the full sample period and for different seasons. In general the autocorrelations are small and statistically insignificant. There is some evidence of small positive autocorrelations at lag 1, which might seem surprising since bid-ask bounce should induce a small negative autocorrelation. Note, however, that returns are computed, in most cases, as average returns across two contracts (see Section 3.3). To the extent that the final trades on these contracts at the end of the day are not perfectly synchronous, this averaging will induce positive serial correlation (Scholes and Williams (1977)). In any case, the autocorrelations provide no evidence of over- or under-reaction. As an alternative, Table 4, Panel C reports future conditional

mean returns, conditional on a large return on day 0. Over- and under-reaction predict return reversals and return continuations, respectively. Again, there is no convincing evidence of either phenomenon. Average returns the first day after a large price move, defined as greater than 2.5% or 5% in absolute value, are small and negative regardless of the direction of the original move. Subsequent returns over the first week, the second week, and the remainder of the month, do not appear to follow any particular pattern. In terms of the magnitudes of these returns, the only notable results are for returns on days 2 to 5 following price increases. These results imply returns over the 4-day period of 1% and 2% for thresholds of 2.5% and 5% for the original price move, respectively. These returns are statistically significant, but given the absence of any obvious pattern and the dangers of searching across numerous different statistics, they are likely the result of sampling variation.

Finally, the theory predicts that the relation between FCOJ futures returns and the temperature should be nonlinear. As the temperature drops further below freezing, the impact on production (duration of the freeze aside) should be more severe. The regression results of Table 5 strongly support the nonlinear nature of the theoretical relation. The coefficient on the nonlinear term in the polynomial regression is highly significant with a t-statistic of approximately 7. Figure 3 graphs the fitted relations of all four models, and both the kernel regression and nonlinear specification show a highly convex relation between FCOJ returns and freezing temperature levels. The piecewise linear model captures the convexity between freezing and above freezing temperature but fails to capture the convexity within the low temperature region. Further subdividing the low temperature range may permit a better fit, i.e., less misspecification, but estimation error would also increase.

4.2 Another Look

The results of Section 4.1 are in stark contrast to the existing view in the literature about weather and FCOJ returns. They show that researchers need to be cautious in how they interpret the relation between asset prices and fundamentals. In this case, the theory predicts a strong relation only under certain conditions, i.e., freezing temperatures. In fact, when those conditions are met, the R^2 s are very high. The existing literature concludes that temperature has limited explanatory power for FCOJ futures returns because it implicitly focuses on the 99% of observations that by theory have no predicted relation. The fact that most of the time there is no relation between

temperature and returns is good news for the theory and market efficiency, not bad news. One should note, however, that the R^2 s are on the order of 50%. This suggests there is substantial variation left to be explained. Of course, the empirical models in Section 4.1 ignore potentially important information and assume that the actual temperature describes the entire shift in the likelihood of there being a freeze and its magnitude. While the true likelihood will be difficult for the researcher to uncover, it is possible to gather more evidence on the true functional form. In particular, beyond the actual realization of a freeze, what other information might be helpful in describing the shift in the freeze likelihood?

4.2.1 Forecasts

One issue, highlighted by Roll, is the fact that the market also has access to temperature forecasts. Roll looks at the relation between returns and temperature surprises (the difference between the realized temperature and the forecast), but he does not focus on these surprises when freezing temperatures are either predicted or realized, i.e., when they will tell us something about the market's change in perception of the likelihood of a freeze. Theory suggests that this is the only time when forecasts should matter. We rectify this omission by using freeze related temperatures and forecasts and by examining R^2 s for different groups of observations as above. The intuition for the role of forecasts is simple. For example, if the market knew with probability one that there would be a freeze the next day, then the actual realization would have no information (ignoring the issue of the severity of the freeze). The price movement due to the fully anticipated freeze would occur the day before. Thus, the model of Section 4.1, which ignores forecasts, provides a lower bound on the true model's ability to explain the relation between FCOJ futures returns and temperature.

There are two primary ways forecasts can impact FCOJ futures returns. First, as described in the extreme example above, forecasts measure the current expectation, though not the entire distribution, of the future temperature and thus contain information about a freeze. Thus, FCOJ futures returns might move prior to the realization of a freeze because the market's estimated likelihood of a freeze has already incorporated the forecast. Second, the actual realization of the temperature might be a surprise relative to the forecast, which also signals a shift in the likelihood (and magnitude) of a freeze. For example, consider the case in which FCOJ returns are negative

when we get positive surprises in the weather. This would be the case when a freeze was forecasted but one did not materialize. This is a case that we do not even try to explain in Section 3.1. Alternatively, there are cases when the freeze was either worse or better than expected relative to the forecast. This too in theory would cause a shift in the distribution of the freeze magnitude and thus FCOJ returns.

Forecasts are only relevant to the extent that the National Weather Service has some forecasting power. Roll (1984) provides a detailed examination of the NWS's forecasting ability, and, in general, finds they produce unbiased forecasts with R^2 s on the order of 50%. We duplicate Roll's tests for a limited 10-year sample of National Weather Service (NWS) temperature forecasts of the next-day minimum temperature for the Orlando area and also look at a time series approach to temperature forecasting. We estimate two basic types of regressions, one including lagged minimum temperatures and forecasts of the form

$$W_t = \alpha + \beta_1 W_{t-1} + \beta_2 W_{t-2} + \gamma F_{t-1}(W_t) + \epsilon_t$$

and one including lagged temperatures and monthly dummy variables of the form

$$W_t = \alpha_i D_{it} + \beta_1 W_{t-1} + \beta_2 W_{t-2} + \epsilon_t,$$

where $F_{t-1}(W_t)$ is the one-day ahead forecast and D_{it} are monthly dummies. We only consider two temperature lags because further lags add little or nothing to the explanatory power of the regressions. Also, we do not include monthly dummies in the forecast regressions because again they do not significantly improve the fit. The estimation results are reported in Table 6. It is important to note that the pure time series regressions are run only on winter season data and that the forecast regressions are for November through March. Variability is much higher in the winter months and forecastability is much lower.

There are several interesting results from the regressions. First, the NWS forecast is not conditionally unbiased since the coefficient is significantly different from 1, but this may simply be due to rounding in high temperature regions. In fact, the NWS often rounds to the nearest 5 degrees when minimum temperatures are high. However, the forecast is quite powerful with an R^2 of 56%. Second, lagged temperatures appear to provide additional forecasting power, increasing the R^2 to 73%, although, again, NWS rounding may be an issue. Minimum temperatures are persistent, and a fall (rise) in minimum temperatures over the previous 2 days suggests a continuing downward

(upward) trend as indicated by the negative coefficient on the second lag, although this coefficient is marginally significant at best. Third, when forecast data is unavailable, monthly temperature dummies are extremely significant, but persistence and trend continuation are still evident. The R^2 is a meaningful 41%. These latter results are consistent with the time series evidence in Campbell and Diebold (2001), who use a much larger sample to forecast temperatures in 10 U.S. cities (although not Orlando).

Unfortunately, the overall ability of the models to forecast temperature is not that relevant. What matters is their ability to forecast temperatures in the freezing range, since that is the only temperature level that affects production and thus FCOJ pricing. The models may be very efficient at predicting 50-degree nights in Florida, but what market participants care about is their ability to predict 25-degree nights. On this issue the pure time series model is not that useful. It never predicts freezes due to the strong reversion to a monthly mean that is well above freezing, even in January. Similarly, the time series augmentation of the NWS forecast does not appear to help in predicting freezes. Perhaps a better way to look at the data is to examine conditional freeze probabilities, which are reported at the bottom of Table 6. The unconditional probability of a freeze is 0.7%. This probability rises to 2.66% in winter and 4.75% in January.

Do NWS forecasts help predict freezes? When the forecast is 32 degrees or below, the probability of observing a freeze the following day is 43.75%. The NWS clearly has some, though clearly imperfect, ability to predict freeze-level temperatures. Lagged temperatures also have predictive power. If the temperature is currently freezing, there is a 32% probability that it will freeze tomorrow as well. Even when it is cold but above freezing, the conditional probability of a freeze of 6.7% is above the January baseline level. Nevertheless, temperature observations do not appear to have much meaningful predictive power after taking NWS forecasts into account.

The NWS's forecasting ability aside, the remainder of this section explores how well forecast information explains FCOJ return variability. We use the following methodology to investigate this issue. First, we generate a set of pricing errors, FCOJ futures return minus the fitted value from the kernel regression of Section 4.1. We then look at whether forecasts can help improve the model, either through forecasting future freezes or through temperature surprises, by regressing the pricing errors on the forecast information. If forecasts are not helpful, then the R^2 s and coefficients should be close to zero. If forecasts are helpful, then we could in principle incorporate them in a

more complete model of FCOJ returns. However, our limited sample of forecast data prevents us from estimating a multi-dimensional nonlinear model with any accuracy.

Table 7 reports results from regressing the pricing errors on information about forecasts through two different sources: (i) forecasts today of freezing temperatures tomorrow, and (ii) realizations of freezing temperatures tomorrow.¹² We also break up the sample into periods in which either a freeze was forecasted or a freeze occurred. That is, we consider both univariate and multivariate versions of the regression,

$$\epsilon_t = \alpha + \beta_1 \text{Max}[0, 32 - F_t(W_{t+1})] + \beta_2 \text{Max}[0, 32 - W_{t+1}] + \nu_t,$$

where ϵ_t is the pricing error from the nonparametric regression, $\text{Max}[0, 32 - F_t(W_{t+1})]$ is tomorrow's forecast of a freeze, and $\text{Max}[0, 32 - W_{t+1}]$ is the realization of a freeze. The table also reports R^2 's calculated within buckets, sorted by the forecast or the future realized temperature.

The overall conclusion from the table is that forecasts have significant but not overwhelming explanatory power for returns, i.e., the relation between FCOJ futures returns and temperature is even stronger than implied by Section 3.1. In the regression of returns on contemporaneous forecasts of tomorrow's minimum temperature, the coefficient is significant and has the correct sign (i.e., forecasts of lower temperatures generate higher positive returns), and the R^2 is 17% in the only relevant bucket (i.e., when the forecast is low). When used on its own, the future temperature has limited explanatory power (e.g., 4% in the relevant temperature range for the regression that uses the full set of winter pre-freeze data) but the correct sign. When both the forecast and future temperature are combined, the R^2 jumps to 24% in the relevant forecast bucket, although the coefficient on the future temperature reverses sign. This sign reversal is puzzling, but it may be attributable to a combination of multicollinearity between the two variables and the relatively small sample size in the relevant region.

As an alternative, Table 8 investigates whether temperature forecast surprises help explain the kernel regression model's pricing errors, and thus potentially improve the model's fit. Because in theory the temperature surprise has a different effect depending on whether a freeze was forecast and/or realized, we consider error analysis regressions of the following sort:

$$\epsilon_t = \alpha + \beta_1 I_{1,t} Z_t + \beta_2 I_{2,t} Z_t + \beta_3 I_{3,t} Z_t + \beta_4 I_{4,t} Z_t + \nu_t$$

¹²The latter piece of information can be broken up into the forecast and the unexpected component of the temperature. We include this latter variable as a noisy signal of the forecast because we have data for the entire sample.

where

$$\begin{aligned}
Z_t &= W_t - F_{t-1}(W_t) \\
I_{1,t} &= 1 \text{ for } F_{t-1}(W_t) \leq 32^\circ \text{ and } W_t > 32^\circ, 0 \text{ otherwise} \\
I_{2,t} &= 1 \text{ for } F_{t-1}(W_t) \leq 32^\circ \text{ and } W_t \leq 32^\circ, 0 \text{ otherwise} \\
I_{3,t} &= 1 \text{ for } F_{t-1}(W_t) > 32^\circ \text{ and } W_t \leq 32^\circ, 0 \text{ otherwise} \\
I_{4,t} &= I_{1,t} + I_{2,t} + I_{3,t}.
\end{aligned}$$

The R^2 s are calculated within buckets, sorted by the forecast or the temperature. Forecast surprises seem to matter under two circumstances—when a freeze was forecast and realized, and when a freeze occurred but was not forecast. The former case generates a negative and significant coefficient (β_2), as predicted by the theory, and is responsible for the 24% R^2 in the low forecast bucket. When the freeze was more severe than predicted, returns are higher. The latter case generates a significant coefficient of the wrong sign and is responsible for the 27% R^2 in the 41 to 45 degree forecast bucket. When the observations are sorted by temperature, the explanatory power of these two regressors is combined in the low temperature bucket, with a resulting R^2 of 31%. Although this estimation is carried out on a limited sample due to the unavailability of extensive forecast data, one could think of extrapolating the results to the full sample. Specifically, the contemporaneous temperature alone explains approximately 50% of the variation of returns in the low temperature region (see Figure 2), and the temperature surprise explains approximately one third of the remaining variation. The combined explanatory power is therefore approximately 65%. In other words, the contemporaneous minimum temperature plus the temperature surprise (relative to the previous day’s forecast) can explain almost two thirds of the return variation in by far the most volatile subset of days, namely when temperatures are low.

4.2.2 What’s Missing?

Realized temperatures alone explain a significant fraction of variation in FCOJ futures returns, and forecasts contribute additional explanatory power. Nevertheless, there is still some unexplained variation in returns in winter months relative to the spring and summer. There are three possible explanations for this phenomenon.

First, the realized minimum temperature may not be a perfect proxy for the severity of the

freeze. Table 1 indicates that freeze duration is also an important variable. In an attempt to address this issue, we analyzed hourly Orlando temperatures, as provided by the National Weather Service, during freeze episodes for the 1983-1998 period. Using these data, we constructed a number of measures of duration based on the critical temperature levels in Table 1. Perhaps unsurprisingly, our duration measures were highly correlated with the minimum temperatures. As a result, given the small sample size, it was impossible to detect a significant effect.

Second, the severity of the freeze may not be a perfect proxy for the resulting supply shock for a number of reasons. For example, the timing of the freeze is important, especially with respect to the EM varieties, which are harvested primarily in January. In our sample, the earliest freeze is in late December and the latest is in late February (see Table 2). Clearly, even controlling for severity, these freezes will not have the same impact on production. There is also the possibility of time dependence. Freezes that occur in years following a freeze year or sequence of freeze years may have different effects on production. In addition, there is a degree of nonstationarity induced by the fact the Florida orange production has gradually been moving southward over the sample period. Orlando temperatures are slowly becoming less relevant, and freezes are generally becoming less of an issue. Tables 2 and 3 provide casual evidence of these combined effects. There is far from a one-to-one correspondence between the minimum temperature reported in Table 2 and the change in the production forecast reported in Table 3. Moreover, looking more closely at the monthly production forecast updates leads to the same conclusion. Ideally, we would relate daily returns to the market's daily unexpected production shock, but this latter variable is unavailable. However, we do look more closely at the monthly production forecast data in the next section.

Finally, even knowing the shock to the market's production forecast is not sufficient. The impact on returns of a given shock to production will depend on a variety of other factors. For example, the shape of the demand curve will influence the price effect of a fixed production shock, given different initial levels of anticipated supply. The availability and price of substitutes (e.g., imports from Brazil, Mexico and the Caribbean) will also influence the price impact. This availability is itself determined by a myriad of other factors, including the weather in other orange growing regions and tariffs, which changed considerably over the sample period. We examine some of these factors in the next section.

Given the complexities outlined above, the explanatory power we document using a simple and

naive model is even more remarkable.

5 Other Components of FCOJ Return Variation

The results of Sections 3 and 4 demonstrate three important results: (i) there is large variation in FCOJ futures returns consistent with theory, i.e., around freezing temperatures, (ii) the relation between FCOJ futures returns and temperature conforms to theory, being nonlinear and convex at temperatures around 32 degrees and below, and (iii) other information, such as forecasts, have an important (though difficult to measure) impact on FCOJ futures returns.

We can use these results to better understand the sources of return variation of FCOJ futures returns. Consider Table 4A, which decomposes the variance of FCOJ returns by season. Approximately 50% of all the variance occurs in winter months, 20% in fall months, and the remaining 30% is shared equally between spring and summer months. Moreover, days with freezing temperatures represent only 4.3% of winter days, but they account for almost two-thirds of winter return variance. In fact, they represent only 1% of all the observations yet capture one-third of all return variance!

Ideally, we would like to be able to explain the remaining return variation, but the thesis of this paper is that, in general, econometricians do not do a good job of modeling complex relations, especially in the absence of strong and reliable theory. Nevertheless, even without knowing the true model, it is possible to show that the underlying economics in the FCOJ futures market may be working rationally with respect to some other important factors. We leave to other research issues related to (i) the microstructure of this market, (ii) the stochastic convenience yield, and (iii) the demand for FCOJ, although these factors may certainly be important.¹³ Instead, we focus on FCOJ-specific supply-based factors, namely (i) news about Florida FCOJ production, and (ii) news about Brazil's FCOJ production. We provide some rudimentary evidence that these factors may also play an important role in FCOJ futures pricing.

Motivated by Roll's (1984) analysis of Wall Street Journal (WSJ) stories, we supplement our analysis by searching the WSJ for articles about FCOJ futures from January 1, 1984 to November 11, 1998. However, in contrast to Roll, who emphasizes that there are many articles about the

¹³Roll examines some of these factors as well as others and finds little or no explanatory power, albeit in a linear framework.

weather, we wish to focus more on other factors. During this period of 3686 days, there are a total of 384 articles in the WSJ about FCOJ. We classify each article into one of six categories related to the focal point of the article, namely news about (i) weather, (ii) production, (iii) Brazil, (iv) demand, (v) technical factors, and (vi) miscellaneous. For example, a news article about Brazil might relate to information about export prices, shipment delays, tariff rates, and Brazilian weather such as droughts. In contrast, demand stories relate to changes in the movement of FCOJ, behavior patterns of consumers, competing products, and retail sales. The results are reported in Tables 9 and 10.

Table 9 shows that on days associated with WSJ news the volatility is substantially higher than on days associated with no WSJ news. In particular, the variance is 3.85 times larger on news days. Furthermore, the direction of the FCOJ price movement coincides with the news content, e.g., 0.97% versus - 1.23% on news about low versus high temperatures, respectively.¹⁴ Most interesting is that of the 384 WSJ stories, less than one-third were about weather, and more than 40% were either about U.S. production or Brazil. Moreover, these stories were associated with return volatilities close to or greater than those of weather-related stories. This is strong anecdotal evidence that other information seems relevant for the FCOJ market.

Table 10 provides the distribution of news stories across months. Unsurprisingly, the weather-related articles generally occurred in winter months. However, stories about Brazilian FCOJ production occur throughout the year, which suggests that important information gets released during spring and summer months. Finally, while the WSJ writes some stories about production in winter and early spring months, October is by far the most important month.¹⁵ As discussed below, October is an important month as it represents the first official government forecast of orange production for the forthcoming season.

¹⁴It should be noted that there is a self-selection problem at work here. On days in which FCOJ futures prices moved, WSJ reporters might look for a story to fit the futures price movement. Alternatively, on days with more interesting stories, the WSJ might ignore news about FCOJ futures. In either case, the reader should be cautious interpreting these results beyond the anecdotal nature of the evidence.

¹⁵Roll also notes stories about USDA crop forecasts but states that they are immaterial for explaining return variability.

5.1 Production Data

Table 4A shows that the variability of FCOJ futures returns in the fall season is about 40% greater than the spring and summer seasons. Why? Is there something special about fall? It clearly cannot be the temperature level as no freeze has actually occurred in the fall season (September through November).¹⁶ However, as mentioned above, an important event occurs every October with respect to orange production—the USDA releases its first production forecast of the season. In fact, the USDA provides a prediction of year-end total orange production on a monthly basis from October to July. Usually in August, the USDA issues a report stating the actual annual production for the previous crop year. At the beginning of each month, the USDA sends out surveys to growers to report their expected production as of the first of the month. Then, the production forecasts are compiled from the survey data and released in the second week of the month. Since October represents the first release of the upcoming orange production forecast, it seems reasonable to assume that it may have special importance; however, other monthly forecast releases may also be relevant.

We collected the monthly USDA production forecast, which is the USDA forecast of the year-end orange production of Florida, for the period 1967-1996. Consequently, we have 30 announcement days in each month, which represent about 5% of the total trading days within a given month. In order for these forecasts to affect returns, in a rational setting, they must contain useful information about future production. One way to assess their information content is to analyze the implied temporal resolution of uncertainty. Table 11, Panel A documents the cross-sectional volatility of USDA production forecasts errors, where the forecast errors are calculated as the percentage difference between the current forecast and final production level. Volatilities are calculated for each month across the 30 years in the sample. Two observations are in order. First, there is significant production uncertainty in October in both freeze and non-freeze years, confirming the evidence from Table 3. Thus, whether a freeze takes place is not the only relevant factor for production and thus pricing. Second, the resolution of uncertainty occurs slowly throughout the year. Little or no relevant information comes out in November or December, but after that there is a noticeable, albeit somewhat uneven, monthly decline in uncertainty. Not surprisingly, a great deal of production information is revealed during January and February in freeze years. In non-freeze

¹⁶This is not quite correct as a mild freeze did occur within our sample on November 25, 1970.

years, the smaller amount of uncertainty is resolved later in the season.

These facts together give support for the existence of important fundamental information other than temperature. One possibility is that this information is actually revealed to the market on USDA announcement dates. However, it is also plausible that the USDA production forecasts lag market information. For example, the State of Florida releases crop damage estimates after significant freezes, which may precede and subsume the information in the subsequent USDA announcement. Table 11, Panel B, examines return volatility on USDA announcement dates. We report the daily standard deviation of returns on these days and the daily standard deviation on the remaining non-announcement days. Freeze days are excluded from the sample since they have already been studied extensively. The results are quite striking and consistent with the theory. The standard deviation of FCOJ futures returns on the October USDA announcement day is 9.3%, which is in the same range as the volatility around freezing temperatures. Incredibly, this one day in October represents over 65% of October return variance, 40% of the total variability in the fall season, and more than 7% of the total return variance. While the results for other months are less impressive, announcement days still coincide with up to 16% of the return variance within the month despite the fact that they only account for 5% of the days. The only exceptions to this pattern are November and December, when it is clear from Panel A that the USDA forecasts contain little or no information.

Consistent with theory, volatility is high on the days with the most news about production. The direction of these FCOJ futures return movements are also consistent with theory. For example, Bauer and Orazem (1994) look at the relation between production surprises and FCOJ futures returns, and find a strong negative correlation. They document, for example, that the October forecast explains over two-thirds of the price movement volatility on that day.¹⁷ Moreover, the above results are consistent with the WSJ news story analysis documented in Tables 9 and 10. Both the volatility and direction of FCOJ futures returns are consistent with the WSJ writing stories about FCOJ in October. As mentioned in an earlier footnote, while there is a sample selection issue, the WSJ analysis, and volatility results here, are consistent with a relation between

¹⁷Using our sample, we also confirm Bauer and Orazem's (1994) findings. It is also possible to improve their model's fit by taking into account nonlinearities similar in spirit to Section 3.1. For brevity, the results are not reported in the tables.

asset prices and fundamentals.

5.2 Brazil

Table 3 documents that the average return for the winter season is related to weather as one might expect. While this result represents only an average finding, the data itself is fairly noisy. What other types of news might affect prices?

Market imperfections aside (e.g., tariffs, import quotas, and transactions costs), the U.S. price of oranges used for FCOJ production should reflect the world price of these oranges. This world price in turn depends on the interaction of the worldwide demand for and supply of oranges. Thus, a downward shift in U.S. supply does not necessarily map one-to-one with an increase in FCOJ prices. Moreover, volatility of FCOJ prices in general will depend on information about aggregate supply and demand beyond the U.S.

Brazil and the U.S. account for over 85% of worldwide production of orange juice, with almost all U.S. production being consumed domestically and Brazil production being exported. The origin of FCOJ production in Brazil dates to the 1960's. The motivation for the development was the big freeze that destroyed a great part of the American groves, namely the 13 million adult orange trees that died in 1962 (see Table 2). Since the U.S. industry did not have enough raw materials to supply both their domestic market and the worldwide market, Brazil filled the gap by accelerating the development of their orange production and processing plants. In fact, by 1984, Brazil was the largest producer of FCOJ worldwide.

Interestingly, similar to the U.S. market and its reliance on the Orlando area, the majority of oranges are produced in one area of Brazil, namely Sao Paolo. However, unlike Orlando's predicament with respect to weather, it is droughts rather than freezes that affect Brazilian production.¹⁸ The drought season in Brazil starts in July and ends in November. Total Brazilian supply is hence known by the winter season in the U.S., but may vary year-to-year. Dry weather may retard or damage buds that bear next season's harvest. Therefore, Brazil's news on production usually starts at the beginning of March.

¹⁸Since Brazilian groves are frost-free, summer droughts are usually the weather risk. In general, the state of Sao Paulo is especially suitable for the production of high quality oranges. It is characterized by an absence of frost, coupled with dry winter and heavy summer precipitation. This annual moisture pattern tends to result in a single blooming period, but does subject FCOJ production to droughts.

Table 12, Panel A reports summary statistics for both the level and change in U.S. production of FCOJ, U.S. imports of Brazilian FCOJ and Brazilian exports of FCOJ over the 1977-1996 period. As seen from the table, the mean levels of production in the U.S. and Brazil are of similar magnitude, with the U.S. importing approximately 30% of its FCOJ. While there has been no marked change in the imports of FCOJ through the years, the level of production in both Brazil and the U.S. has drifted up. To understand the implications of these results more fully, consider the following facts about FCOJ: (i) the U.S. is a net importer of FCOJ, (ii) the worldwide supply curve of oranges in any given year is fairly inelastic (i.e., the production levels are fixed, irrespective of price), and (iii) the U.S. demand curve has a higher slope than worldwide demand because, at the very least, it is a subset of that demand.

Given these facts, there are two important implications for the interaction of U.S. production and Brazilian imports and exports. First, relative to what imports would have been without a freeze, one would expect imports to go up (down) given a negative (positive) supply shock in the U.S. Second, and similarly, one would expect imports to go up (down) given a positive (negative) supply shock in Brazil. As a first pass, and making the strong assumption that the previous year's level represents our expectations of future imports, we estimated the correlation between import changes of FCOJ and both changes in U.S. production and Brazil exports. As predicted, the correlations are -0.36 and 0.56, respectively.

Table 12, Panel B looks more closely at the interaction between U.S. imports and supply shocks in the U.S. and Brazil. In particular, we report the magnitude of the average change in U.S. imports of FCOJ conditional on either negative or positive supply shocks to both U.S. and Brazilian production. The results suggest that, for negative supply shocks, U.S. and Brazilian FCOJ are close substitutes to the extent their magnitudes are similar. For example, conditional on negative production in the U.S., the average drop in production is 23.91 million gallons (42 degrees brix) while the corresponding imports of Brazilian FCOJ are 19.26 million gallons. Similarly, when Brazilian production falls, the average drop is 25.34 million gallons with a corresponding 16.83 million gallon drop in U.S. imports. While a similar pattern emerges for positive supply shocks, there is less of an impact on imports, suggesting some nonlinearity in the demand curve or fixed transactions costs. As a final analysis, we regress the change in imports jointly on both the change in U.S. and Brazilian production levels. As expected, the coefficients are negative for U.S. shocks

and positive for Brazil shocks. Moreover, the estimates are both statistically and economically significant. In fact, in terms of explained variation of import changes, these two components alone explain 44%.

While this analysis is anecdotal in nature, it does illustrate two important facts: (i) Brazilian imports are significant in magnitude, and (ii) these imports are used as substitutes for domestic production, especially during low production years (i.e., freeze years). To the extent that there is information about Brazil's FCOJ production, such as transport costs, tariffs, Sao Paulo weather, and, more directly, Brazil's FCOJ production and prices, this should get incorporated into FCOJ futures prices. Thus, this factor represents another important component, not related to the temperature level in Orlando. As supporting evidence, the WSJ news story analysis in Table 9 shows that approximately 15% of all FCOJ stories are related to Brazil, and that on these days the variance of FCOJ returns is more than three times that on non-news days.

6 Concluding Remarks

This paper attempts to reverse the literature's view that FCOJ futures returns are not explained by fundamentals, particularly temperature. We present the following empirical facts that are strongly supported by theory:

- Though covering only 4.5% of the days in winter, two-thirds of all winter return variance in FCOJ futures coincide with freezing temperatures.
- On these days, using a simple, but theoretically appropriate nonlinear model of the relation between FCOJ returns and weather, we can explain almost 50% of the variance. Moreover, these R^2 s can increase to almost 65% if the model is extended to include market information about forecasts.
- Other factors also play an important role in explaining FCOJ price variability. For example, we document that over 40% of the fall season variance of FCOJ futures returns occurs on 1.5% of the days. These days are easy to identify; they are the date of the first USDA forecast of orange production for the upcoming season, released in mid-October.

The bottom line from this paper is that when we have an understanding of the fundamental information, both the FCOJ futures return variability and direction seems to coincide with theory. Thus, this research suggests a strong link between FCOJ futures returns and their underlying fundamentals.

In our opinion, it was this apparent lack of a strong link, as suggested by a low R^2 , that made the FCOJ market such a prominently cited example of excess volatility. Now that we have successfully addressed this issue, it will be tempting for behavioral economists to argue that the real volatility puzzle for FCOJ returns is volatility during times other than freezes. Of course, all assets - other commodities, equities, and bonds, among others - exhibit some volatility throughout the year. Nevertheless, why is there volatility in FCOJ futures returns on days other than around low temperatures or USDA production news? The pricing of FCOJ futures is, like any asset, extremely complicated. The futures convenience yield, market microstructure, long-run news about weather, news about orange quality, news events about Brazil, and short-run demand shifts, among many other factors, can all have important effects. Cursory evidence suggests that this is the case. For example, on news days about Brazilian production, which tend to be in summer, there is substantially higher volatility in the FCOJ futures market. If an econometrician could build the correct structural model with all the relevant interactions and accurately measure the fundamentals, the vast majority of variability might be explained, just as we have done here with respect to temperature. However, this is a difficult, if not impossible, task. Even with respect to FCOJ futures returns and temperature, the relation is nonlinear, multi-dimensional and state and path dependent.

Appendix: Return Calculation - Discussion of Limit Moves

In general we want to focus on the most liquid contracts, which are usually the ones closest to maturity except when close to expiration. To that end we develop an algorithm that results in a spliced return series which utilizes the most liquid contracts. Specifically, we make a distinction between two types of periods: (i) trading days in an expiration month prior to expiration (we denote this period EXP), and (ii) the rest of the time, i.e., trading days in months in which no contracts expire and trading days in an expiration month and past the expiration date of the contract (we denote this period NoEXP). In period NoEXP we want to focus on the two near maturity contracts C1 and C2:

- If neither contract hits its price limit, the return is the average return on C1 and C2.
- If C2 hits the limit and C1 does not, we calculate the return from C1 only.
- If C1 hits the limit and C2 does not, we average the returns on C1 and C2.
- If both contracts limit, we average the returns on C1 and C2 return and aggregate forward until either is off-limit (and record the minimum temperature for the trading “day” as the minimum temperature for the entire limiting period).

In period EXP we want to avoid using C1 due to the illiquidity that develops in futures contracts close to expiration. Our rule above is then applied to C2 and C3 in a similar manner.¹⁹

¹⁹During the sample period a number of rules relevant for the calculation of returns were changed. First, with respect to contract expiration dates, prior to November 1994 the expiration date was 10 business days prior to the end of every other month. In the following period expiration was 14 business days prior to the end of the month. As for the size of the price limits, initially it was 3 cents, and then it moved up to 5 cents in January 1979. From May 1986 there is no price limit on C1. From January 1993 there is no limit on the move in C2 during the expiration month of C1.

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Critical Temperatures for Oranges		
	Winter (Dec to mid-Feb)	Spring (mid-Feb to Mar)
Bloom	NA	29-30
Fruit	26	29
Leaves	24	27
Twigs	22	24
Branches	20	22

Severity of Freeze (Damage vs. Time)		
Temperature	Duration	Consequence
28	6 hours or more	Extensive Fruit Damage
26	4 hours or more	Extensive Fruit Damage
24	2 hours or more	Extensive Fruit Damage
20	4 hours or more	Extensive Tree Damage

Freeze Recovery Periods		
Freeze Type	Minimum Temperature	No. of Years to Recovery
Light	32-29	0
Moderate	28-25	0
Hard	24-21	1-3
Severe	Less than 20	3-7

Table 1: Temperatures and Orange Production Damage

Critical temperatures and durations and freeze recovery periods for orange trees. Source: Citrus Associates of New York Cotton Exchange, Inc, 1994

Date	Low Temperature (Orlando area)	Damage
February 7-9, 1835	11°	Catastrophic
1857	NA	Moderate
December 1, 1876	NA	Moderate/Severe
January 9-13, 1886	19°	Severe
December 29-30, 1894	18°	Severe
February 8-9, 1895	17°	Catastrophic
January 28-29, 1897	25°	Moderate
January 2-3, 1898	23°	Severe
February 13-14, 1899	20°	Catastrophic
January 1900	27°	Moderate
February 17-18, 1900	26°	Moderate
December 21, 1901	26°	Moderate/Severe
January 14, 1902	26°	Severe
January 26-27, 1905	21°	Severe
February 2-6, 1917	22°	Severe
December 12-13, 1934	22°	Severe
January 27-29, 1940	22°	Severe
February 5-6& 11, 1947	24°	Moderate/Severe
December 12-13, 1957	21°	Moderate/Severe
January 9-10, 1958	26°	Severe
February 4-5, 14, 17-19, 21, 1958	24° – 26°	Severe
December 12-13, 1962	19°	Severe/Catastrophic
January 31, 1966	23°	Moderate
January 8-11, 1970	24°	Moderate/Severe
January 20-22, 1971	22°	Severe
January 19, 1977	20°	Severe
January 12-14, 1981	18°	Severe
January 11-12, 1982	23°	Severe
December 24-25, 1983	24°	Severe
January 20-22, 1985	19°	Severe*
December 26-27, 1985	26°	Moderate
January 28, 1986	27°	Moderate
February 24, 1989	29°	Moderate
December 23-24, 1989	25°	Severe
January 8-9, 1996	29°	Moderate
February 5-6, 1996	26°	Moderate

Table 2: History of Freezes

A list of the important freezes from 1835 to 1998 as documented by Attaway (1997). An important freeze is one in which the temperature in the Orlando area (i) dropped substantially below freezing, and (ii) remained at this level for a significant period of time. The description of the freezes, varying from moderate to catastrophic, is subjective, based on a description of the economic impact of the freezes provided in Attaway (1997). Note that the * refers to the fact that, although each individual freeze in the 1981-85 period was severe, the combined effect of four freezes in five years was generally considered catastrophic. Our sample period is 1967-98 and therefore covers the freezes below the horizontal line.

Year	Freeze	Production Change (%)	Cumulative Return(%)
67		4.7	-12.6
68		4.6	31.1
69	Yes	-3.7	-8.4
70	Yes	-15.3	16.2
71		5.0	-14.6
72		-2.5	-6.2
73		3.6	-16.1
74		-0.4	-19.5
75		5.3	-2.1
76	Yes	-10.6	52.7
77		2.3	-0.7
78		-1.8	-7.8
79		3.3	-12.2
80	Yes	-15.1	44.5
81	Yes	-24.2	-0.6
82		-2.4	-12.6
83	Yes	-30.5	30.1
84	Yes	-12.7	-3.3
85	Yes	-9.7	-27.7
86		-7.2	0.2
87		6.2	-4.1
88	Yes	-3.6	-14.9
89	Yes	-15.2	55.0
90		-8.1	1.6
91		2.8	-16.1
92		0.3	-34.3
93		1.4	7.1
94		4.8	-9.8
95	Yes	0.6	-3.5
96		2.8	-23.6
Frz		-12.7	12.7
Non-Frz		1.3	-6.1

Table 3: Freezes, Production and Returns

Year by year percentage differences between actual orange production and the October USDA forecast at the beginning of the season, and the cumulative FCOJ futures return over the corresponding winter season (December to February). The bottom two rows show the averages of these numbers over freeze years and non-freeze years. Seasons are classified as freeze/non-freeze based on Attaway (1997).

Panel A: Seasonal Returns							
Season	Obs	Mean (%)	SD (%)	AC(1)	AC(2-20)	% Obs	% Var
All	7543	0.033	2.105	0.018	0.001	100.0	100.0
Winter	1772	0.009	3.160	0.029	0.001	23.5	52.9
Pre-Freeze	1475	0.015	3.305	0.005	-0.005	19.6	48.2
Post-Freeze	297	-0.020	2.315	0.055	0.008	3.9	4.8
Fall	1878	0.068	1.847	-0.016	0.005	24.9	19.2
Spring	1939	0.059	1.555	0.030	-0.006	25.7	14.0
Summer	1954	-0.002	1.539	0.009	-0.007	25.9	13.9

Panel B: Temperature-Sorted Returns (Winter Pre-Freeze)					
From-To	Obs	Mean (%)	SD (%)	% Obs	% Var
0-99	1475	0.015	3.305	100.0	100.0
0-35	77	2.721	11.815	5.2	69.4
36-40	134	-0.311	2.163	9.1	4.0
41-45	198	-0.273	2.081	13.4	5.4
46-99	1066	-0.086	1.790	72.3	21.3

Panel C: Conditional Mean Returns						
Threshold	Obs	Horizon (Days)				
		1	2-5	6-10	11-20	
<-5.0%	69	-0.076	-0.112	0.111	0.044	
<-2.5%	419	-0.051	0.076	0.154	0.065	
>2.5%	374	-0.061	0.256	0.039	0.044	
>5.0%	74	-0.047	0.527	0.039	0.011	

Table 4: FCOJ Futures Returns

Panel A presents means, standard deviations and autocorrelations (at lag 1 and the average over lags 2 to 20) of daily FCOJ futures returns in percent for different seasons, and the percentage of the variance of returns accounted for by the observations in each season. Panel B presents the same data (with the exception of autocorrelations) for different temperature ranges. Panel C presents mean daily returns over different horizons conditional on the level of the return on day 0. Continuous futures daily return series are constructed off of the three closest maturity contracts in a manner described in Appendix 1. The data period is September 1967 to August 1998. Temperatures are the minimum temperature for the Orlando region, contemporaneous with the close-to-close futures return. Winter season is defined as December, January and February; Spring is March, April and May; Summer is June, July and August; and Fall is September, October and November. The pre-freeze period includes days up to and including the first freeze of the season, if a freeze occurs.

	α	β	β_1	β_2	R^2
Winter Pre-Freeze					
Linear	2.187 (0.474)	-0.042 (0.009)			0.015
Piecewise					0.281
0 – 35	58.113 (8.409)	-1.741 (0.262)			
36 – 40	7.714 (5.419)	-0.211 (0.142)			
41 – 45	6.620 (4.590)	-0.159 (0.106)			
46 – 99	-0.558 (0.499)	0.008 (0.009)			
Nonlinear	-0.135 (0.071)		0.095 (0.339)	0.202 (0.031)	0.333
Kernel					0.336
Full Sample					
Nonlinear	0.003 (0.022)		-0.509 (0.187)	0.253 (0.018)	0.162
	-0.006 (0.022)		1.971 (0.056)		0.141

Table 5: Model Estimates and R^2 s

Estimation results for the four models of FCOJ futures returns as a function of contemporaneous minimum temperatures as described in Section 3.1. The data period is September 1967 to August 1998. The winter, pre-freeze period includes days in December, January and February up to and including the first freeze of the season, if a freeze occurs. Continuous futures return series are constructed off of the three closest maturity contracts in a manner described in Appendix 1. Returns are in percent per day. Temperatures are the minimum daily temperature for the Orlando region, contemporaneous with the close-to-close futures return.

	Const.	W_{t-1}	W_{t-2}	$F_{t-1}(W_t)$	R^2				
W_t	13.96			0.74	0.56				
	(1.22)			(0.02)	853				
W_t	18.11	0.70	-0.04		0.46				
	(1.49)	(0.03)	(0.03)		853				
W_t	0.10	0.46	-0.03	0.56	0.73				
	(1.23)	(0.03)	(0.02)	(0.02)	853				
	W_{t-1}	W_{t-2}	Dec.	Jan.	Feb.	R^2			
W_t	0.67	-0.07	21.05	19.77	20.51	0.41			
	(0.02)	(0.02)	(1.03)	(1.05)	(1.08)	1877			
	All	Winter	Dec	Jan	Feb	Forecast	Temperature		
						0-32	33-40	0-32	33-40
Prob (%)	0.70	2.66	1.57	4.75	1.53	43.75	4.29	32.00	6.69
Obs	7756	1877	636	653	588	16	70	50	254

Table 6: Temperature Forecasting

Estimation results for the temperature forecasting models as described in Section 3.2.1 and conditional freeze probabilities. Conditional probabilities are the probability of observing a minimum temperature less than or equal to 32° , conditional on the season/month, previous day's forecast, and previous day's temperature. Temperatures are the minimum daily temperature for the Orlando region and forecasts are one-day ahead minimum temperature forecasts. Standard errors are in parentheses under the coefficient estimates, and the number of observations in each regression appear under the regression R^2 s.

α	β_1	β_2	Bucket	R_{0-35}^2	R_{36-40}^2	R_{41-45}^2	R_{45-99}^2	R_{0-99}^2
-0.09 (0.11)	0.24 (0.15)	-	$F_t(W_{t+1})$	0.17 26	0.00 38	0.00 70	0.00 309	0.01 443
-0.07 (0.11)	-	0.02 (0.11)	W_{t+1}	0.00 21	0.00 30	0.00 60	0.00 332	0.00 443
-0.09 (0.11)	0.52 (0.23)	-0.26 (0.17)	$F_t(W_{t+1})$	0.24 26	0.00 38	0.03 70	0.00 309	0.01 443
			W_{t+1}	0.02 21	0.00 30	0.05 60	0.01 332	0.01 443
0.01 (0.06)	-	0.12 (0.07)	W_{t+1}	0.04 79	0.00 132	0.00 197	0.00 1056	0.00 1464

Table 7: Error Analysis - Forecasts

Estimation results from error analysis regressions, where the errors are from a winter pre-freeze kernel regression of returns on contemporaneous minimum temperatures. The model, described in Section 3.2.1, has contemporaneous forecasts and future temperature as explanatory variables. The R^2 is calculated within buckets, sorted by the forecast or the temperature, as denoted in column “Bucket”. Data is for days where forecasts are available, except the final regression, which is for all winter pre-freeze days except freeze days (hence 1464 days: 1475 winter pre-freeze observations less 11 freeze days in our sample). Standard errors are in parentheses under the regression estimates, and the number of observations in each bucket appear under the bucket R^2 s.

α	β_1	β_2	β_3	β_4	Bucket	R^2_{0-35}	R^2_{36-40}	R^2_{41-45}	R^2_{45-99}	R^2_{0-99}
-0.07	0.00	-1.58	0.48	-	$F_{t-1}(W_t)$	0.24	-0.01	0.27	0.00	0.11
(0.11)	(0.06)	(0.27)	(0.11)			26	37	70	310	443
					W_t	0.31	0.00	0.00	0.00	0.11
						20	30	59	334	443
-0.07	-	-	-	0.05	$F_{t-1}(W_t)$	-0.02	0.00	0.05	0.00	0.00
(0.11)				(0.06)		26	37	70	310	443
					W_t	0.01	0.00	-0.01	0.00	0.00
						20	30	59	334	443

Table 8: Error Analysis - Temperature Surprises

Estimation results from error analysis regressions, where the errors are from a winter pre-freeze kernel regression of returns on contemporaneous minimum temperatures. The model, described in Section 3.2.1, has contemporaneous temperature surprises (temperature less the prior day’s forecast) as the explanatory variable, grouped by the level of the temperature and the forecast. The R^2 is calculated within a bucket, sorted by the forecast or the temperature, as denoted in column “Bucket”. Data is for days where forecasts are available, for a total of 445 days. Standard errors in parenthesis under regression estimates, and number of observations in each bucket appear under bucket R^2 s.

	Sample Size	Mean (%)	SD (%)
News Day	384	0.015	3.06
No News Day	3302	-0.008	1.55
Weather	121	0.320	3.30
Production	101	-0.034	3.85
Brazil	58	0.554	2.95
Demand	21	0.065	2.90
Technical	50	0.307	2.80
Miscellaneous	33	-0.042	1.27
Weather (low)	73	0.966	3.112
(high)	48	-1.230	2.637
Production (low)	46	0.928	3.534
(high)	55	-0.869	2.437

Table 9: FCOJ Futures Returns and WSJ Events

The means and standard deviations of daily FCOJ futures returns on all the days the Wall Street Journal (WSJ) reported some story about FCOJ futures. In particular, the sample covers all WSJ articles from January 1, 1984 to November 11, 1998 in which FCOJ futures contracts were discussed. This sample represents 3,686 trading days. Based on the content of the article, the stories were broken down into one of six categories related to a particular fundamental factor: (i) weather, (ii) production, (iii) Brazil, (iv) demand, (v) technical, or (vi) miscellaneous.

	Oct	Nov	Dec	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sept
Weather	4	12	41	45	21	7	0	0	1	0	2	0
Production	31	1	2	12	18	12	11	3	3	1	6	6
Brazil	6	3	3	9	3	9	3	6	9	4	3	2
Demand	0	2	3	0	0	3	1	4	1	3	3	1
Technical	3	7	3	3	3	5	8	3	6	7	2	4
Misc.	4	1	0	3	3	0	3	0	3	5	6	7

Table 10: Number of News Stories by Month

Distribution of news stories by month. News stories related to FCOJ as reported in the Wall Street Journal from 1984-1998. The stories are broken down into one of six categories: (i) weather, (ii) production, (iii) Brazil, (iv) demand, (v) technical and (vi) miscellaneous.

Panel A: Production Uncertainty										
	Oct	Nov	Dec	Jan	Feb	Mar	Apr	May	Jun	Jul
Frz	12.9	12.9	13.6	11.9	6.1	3.5	3.7	2.4	1.8	0.4
Non-Frz	4.2	4.1	4.3	3.8	3.8	2.8	2.4	1.8	1.2	0.5
Panel B: Announcement Day Volatility										
	Oct	Nov	Dec	Jan	Feb	Mar	Apr	May	Jun	Jul
Ann.	9.258	1.542	2.530	2.488	3.046	2.190	2.294	2.790	2.139	1.670
Non-Ann.	1.464	1.510	2.531	2.315	1.651	1.651	1.372	1.452	1.455	1.597
% Ann. Days	4.6	5.0	4.9	4.9	5.3	4.6	4.7	4.6	4.6	4.7
% Variance	65.3	5.1	4.8	5.5	16.0	7.9	12.6	15.0	9.2	5.2

Table 11: Production Uncertainty

Panel A reports month by month volatility of the USDA production forecasts errors. Forecast errors are calculated as the percentage difference between the forecast in a given month and the final production. The table reports the cross-sectional (across years) standard deviation of these forecast errors. Panel B reports return volatility on the announcement day for each month, the volatility for others days in the month, the percentage of days which are announcement days, and the percentage of variance occurring on announcement days (freeze days are excluded from the analysis).

Panel A: Descriptive Statistics						
	U.S Production	Brazil Exports	U.S. Imports from Brazil	Change in U.S Production	Change in Brazil Exports	Change in U.S. Imports from Brazil
Mean	237.3	253.8	99.0	6.7	17.0	1.3
Std. Dev.	33.9	102.1	38.1	31.7	54.4	29.5
Corr. with Δ Imp.				-0.36	0.56	

Panel B: Explaining Imports			
State	Average Δ U.S. Production	Δ Brazil Exports	Δ U.S. Imports
Δ U.S. Prod.>0	24.5		-9.1
Δ U.S. Prod.<0	-23.9		19.3
Δ Brazil Exp.>0		64.0	21.5
Δ Brazil Exp<0		-25.3	-16.8

Regression Results				
	Const.	Δ U.S. Prod.	Δ Brazil Exp.	R^2
Δ U.S. Imp.	-1.56	-0.34	0.30	44.4
	(5.74)	(0.17)	(0.10)	

Table 12: U.S. and Brazilian FCOJ Production

Panel A reports means, standard deviations and correlations for the level and annual change in U.S. production, Brazilian exports, and U.S. imports from Brazil of FCOJ for the sample period 1977-96. All quantities are in millions of gallons (42 degrees brix). Panel B reports average annual changes in U.S. production, Brazilian exports, and U.S. imports from Brazil of FCOJ conditional on positive and negative changes in the former two variables. It also reports results from a regression of changes in imports on changes in U.S. production and Brazilian exports.

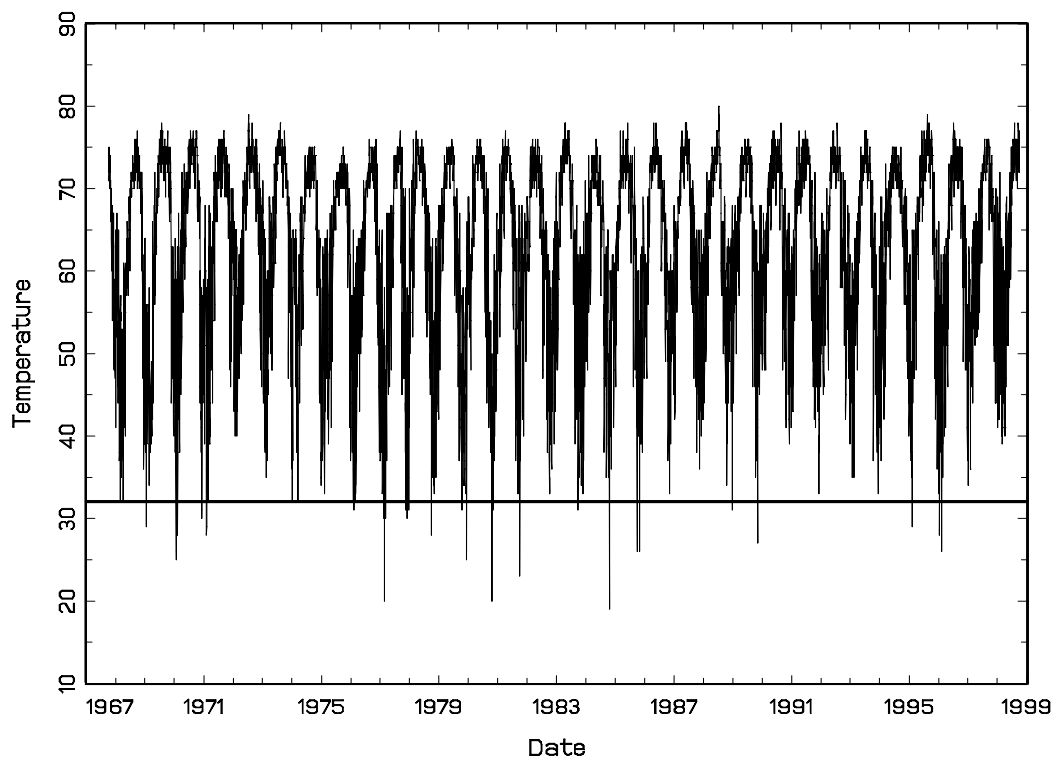


Figure 1: Minimum Temperatures
Daily minimum temperatures over the sample period, September 1967 to August 1998. The Horizontal line represents 32° .

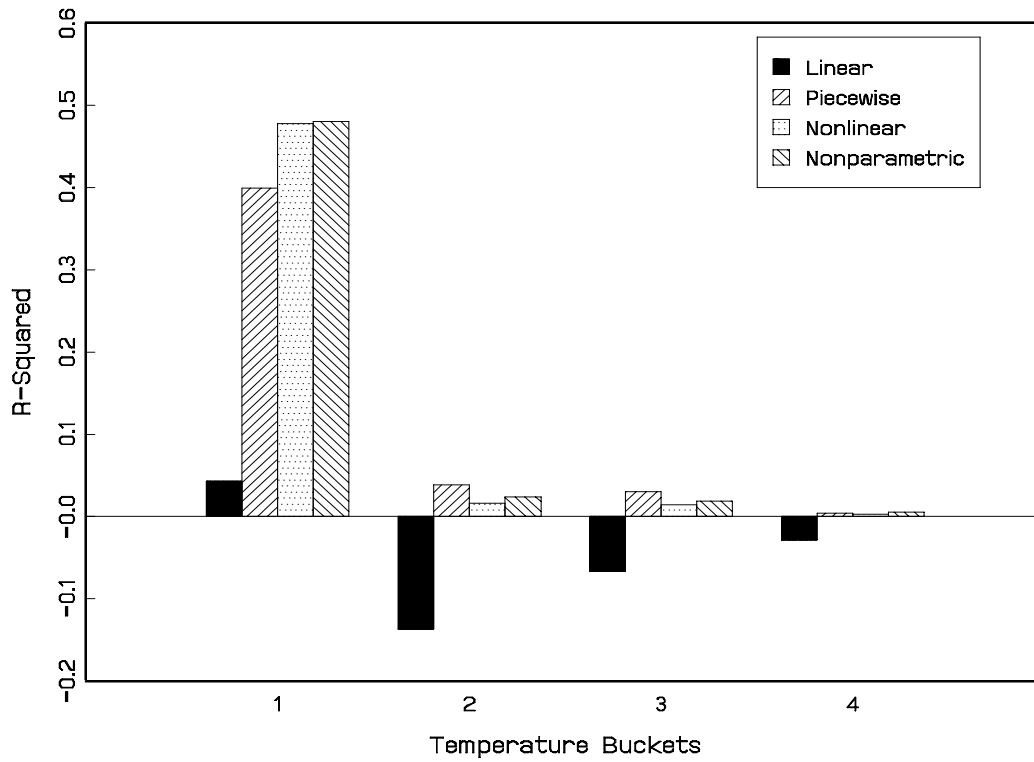


Figure 2: R^2 s by Temperature Bucket

R^2 s by temperature buckets for the four models described in Section 3.1. The temperature buckets are ≤ 35 , 36-40, 41-45 and ≥ 46 . The R^2 s refer to models performed in winter months, pre-freeze.

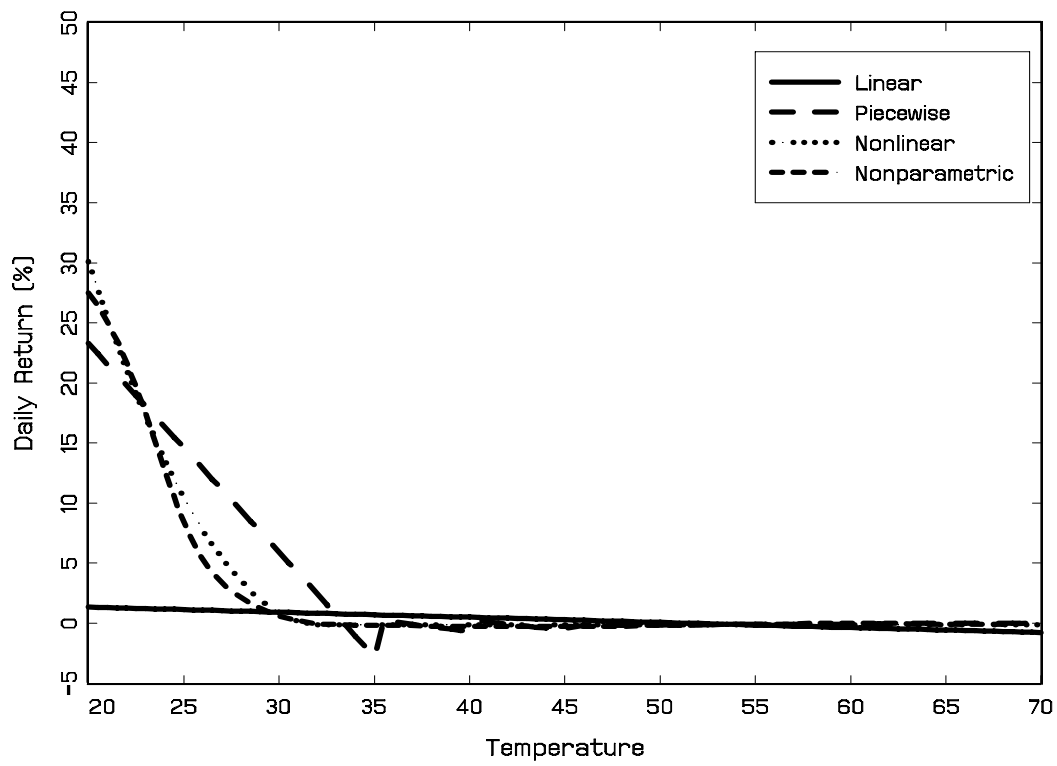


Figure 3: Fitted Returns

Fitted daily returns versus contemporaneous realized temperature for the four models described in Section 3.1. The analysis was performed in winter months, pre-freeze.